## CPS 5310 Spring 2019 Homework and Lab Assignments

Homework 1 (Due Thursday, 1/31)

Refer to the Introduction to Probability Models book (10th Edition) by Sheldon Ross.

(a) Practice:

Chapter 1 #1, 2, 3, 4, 11 Chapter 2 # 1, 3, 5, 7

(b) Write up solutions for the following problems to be turned in on the due date: Chapter 1 #33 Chapter 2, # 2, 4, 12, 16

## Homework 2 (Due Thursday, 2/7)

- (a) Read Chapter 2 of the *Statistics* book by Agresti and Franklin. In particular, make sure you get the meanings of the following terms for a set of numerical data: mean, variance, standard deviation, median, quartiles, five number summary, interquartile range, outlier. Also learn how to visualize data with dot plots, stem and leaf plots, pie charts, box plots, and histograms.
- (b) Practice:

Section 2.1 # 2.6 Section 2.4 #2.49 Section 2.5 #2.74 Section 2.6 #2.89 Chapter Problems #2.116

(c) Write up solutions for the problems below and turn in on the due date:

- Section 2.2 #2.16 Section 2.3 #2.45 Section 2.4 #2.48 Section 2.5 #2.70
- (d) Learn a few basic commands in R to create a data vector or matrix, to read a data file, to perform basic arithmetic operations, to get descriptive summaries and graphical visualizations like those listed in part (a). [Note: Your TA, Ms. Sumi Dey, will cover these during the lab on Tuesday, February 5. Please install R-Studio on your own laptop and bring it to the lab. Or you can use R-Studio on the lab computers.]
- (e) Work out problem #2.117 using R and put your answers in a single document. The data is in the file central\_park\_yearly\_temps.txt. Do NOT mix your answers with the R commands and outputs, which should be put in an appendix at the end of the document. Turn in a copy by email to Sumi (sdey2@miners.utep.edu) on the due date.

Homework 3 (Due Thursday, 2/14)

Refer to the Introduction to Probability Models book by Sheldon Ross.

(a) Practice:

Chapter 2 # 14, 22, 23, 33, 38

(b) Write up solutions for the following problems to be turned in on the due date: Chapter 2, # 34, 40, 49, 52, 53

Homework 4 (Due Thursday, 2/21)

Refer to the Introduction to Probability Models book by Sheldon Ross.

(a) Practice:

Chapter 2 # 29, 36, 41, 42, 50

(b) Write up solutions for the following problems to be turned in on the due date: Chapter 2, # 37, 45, 51, 54, 73

# Solutions to Homework 3 Due to Panfeng Liang

Hw3 panfong liang 34. (a) By properties of cumulative probability  $F(\infty) = \int_{-\infty}^{\infty} f(x) dx = 1$  $\int_{0}^{2} C(4x - 2x^{2}) dx = C \left[ 2x^{2} - \frac{2}{3} x^{3} \right]_{0}^{2}$  $2x4 - \frac{2}{3}x8$  $=\frac{8}{3}C=1$ C= 3/8 Thus, (b)  $p(\frac{1}{2} \le x \le \frac{3}{2}) = \frac{3}{8} \int_{\frac{1}{2}}^{\frac{3}{2}} (\frac{1}{2}x - \frac{1}{2}x^{2}) dx$  $=\frac{3}{8}(2x^{2}-\frac{2}{3}x^{3})\frac{3}{2}$ 0.6875 = 16

40, The distribution table can be shown as below: # of Games probability pt+ (1-p)t 4 (4) P(1-p)p+ (4) (1-p) p(1-p) 5  $(\frac{5}{3})\vec{p}(-\vec{p})\vec{p} + (\frac{5}{3})(1-\vec{p})\vec{p}(1-\vec{p})$  $(\frac{1}{3})p^{3}(1-p^{3}p+(\frac{1}{3})(1-p^{3}p^{3}(1-p))$ Expected value =  $4\left[p^{4}+(1-p)^{4}\right] + \sum \left[p^{4}(1-p)+(1-p)^{4}p\right]$ + 60[p#1-px+(1-p)\*p2]+ 40[p4(1-p)+(1-p)\*p3] p/uq in p=1/2, E=5.8125

49.  $E(x^{2}) - [E(x)]^{2} = var(x) > 0$ when variance = 0, for example, x is a constant, we have equality.

52 (a)  $F_{M}(x) = x^{n}$  $f_m(x) = n x^{n-1}$  $E(M) = \int_{0}^{1} x f_{m}(x) dx = \int_{0}^{1} n x dx = \frac{n x}{n+1}$ = 11 b)  $\int C(1-x^2) dx$  $= C(X - \frac{3}{3}) \Big|_{-1}^{1} = C\Big[(1 - \frac{1}{3}) - (-1 + \frac{1}{3})\Big] = 1$  $C = \frac{2}{4}$  $E(x) = \left( \begin{array}{c} -\frac{3}{4} \times (1-x^{2}) dx = \frac{3}{4} \left( \begin{array}{c} x^{2} - \frac{x^{4}}{4} \right) \right)^{-1}$ (c)  $E(x) = \int_{0}^{1} C x (4x - 2x^{2}) dx = c \int_{0}^{1} (4x^{2} - 2x^{2}) dx$  $= C[4x^{2} - x^{4}] \Big|_{0}^{2} = {}_{3}C(\frac{32}{3} - 8)$  $C=\frac{2}{8}, E(x)=1$ 

B, X~ U(0,1) f(x) = 51,  $x \in (0, 1)$ 0, otherwise  $E(X^{n}) = \int_{0}^{1} x^{n} \cdot 1 \, dx = \frac{n!}{n!} = \frac{1}{0}$  $E(x^{n}) = \int_{0}^{1} x^{n} dx = \frac{2n+1}{2n+1} = \frac{1}{2n+1}$  $var(x^n) = E(x^n) - [E(x^n)]^d$  $=\frac{1}{2n+1}-\frac{1}{(n+1)^2}$ 

# Solutions to Homework 4 Due to Yi Xie



February 15, 2019

# Q37

Let  $X_1, X_2, ..., X_n$  be independent random variables, each having a uniform distribution over (0,1). Let M = maximum  $(X_1, X_2, ..., X_n)$ . Show that the distribution function of M,  $F_M(\cdot)$ , is given by  $F_M(x) = x^n$ ,  $0 \le x \le 1$ 

What is the probability density function of *M*?

## Answer

Since  $X_1, X_2, ..., X_n$  be independent random variables, each having a uniform distribution over (0,1), so we have  $F_{X_1}(x) = P(X_1 \le x) = F_{X_2}(x) = P(X_2 \le x) = \cdots = F_{X_n}(x) = P(X_n \le x) = x$ .

 $P(M \le x) = P(X_1 \le x, X_2 \le x, \dots, X_n \le x) = P(X_1 \le x) \cdot P(X_2 \le x) \dots P(X_n \le x) = x^n \ 0 \le x \le 1.$ 

The probability density function of *M* is  $f_M(x) = \frac{d}{dx}F_M(X) = \frac{d}{dx}x^n = nx^{n-1}$ , o < n < 1

# Q45

A total of r keys are to be put, one at a time, in k boxes, with each key independently being put in box i with probability  $p_i$ ,  $\sum_{i=1}^k p_i = 1$ . Each time a key is put in a nonempty box, we say that a collision occurs. Find the expected number of collisions.

## Answer

Let  $N_i$  denote the number of keys in the i-th box and  $\sum_{i=1}^{n} N_i = r$ . Let I be a indicator function

$$I(N_i = 0) = \begin{cases} 1, \text{ if } N_i = 0\\ 0, \text{ otherwise} \end{cases}$$

Then if  $N_i \ge 1$ , there are  $N_i - 1 + I(N_i = 0) = N_i - 1$  times collisions in the i-th box, if  $N_i = 0$ , then there is  $N_i - 1 + I(N_i = 0) = 0 - 1 + 1 = 0$  collision in the i-th box.

So in the i-th box there are  $N_i - 1 + I(N_i = 0)$  collisions.  $X = \sum_{i=1}^{k} (N_i - 1 + I(N_i = 0))$  is the total collisions in all the box. Take expectation on both side, then we have  $E(X) = E[\sum_{i=1}^{k} (N_i - 1 + I(N_i = 0))] = E[\sum_{i=1}^{k} N_i] - \sum_{i=1}^{k} E[1] + \sum_{i=1}^{k} E[I(N_i = 0)] = r - k + \sum_{i=1}^{k} E[I(N_i = 0)].$  $E[I(N_i = 0)] = 1 \cdot P(I(N_i = 0) = 1) + 0 \cdot P(I(N_i = 0) = 0) = 1 \cdot (1 - p_i)^r$ , so  $E[X] = r - k + \sum_{i=1}^{k} (1 - p_i)^r$ .

### Q51

A coin, having probability p of landing heads, is flipped until a head appears for the r th time. Let N denote the number of flips required. Calculate E[N]. Hint: There is an easy way of doing this. It involves writing N as the sum of r geometric random variables.

#### Answer

The Let experiment A = flipping a coin until a head appears for the r th time. Let experiment B = flipping a coin until we get a head, then experiment A = do experiment B r times.

Let  $N_i$  denote the number of flips required to get a head i = 1, 2, 3, ..., r. Then  $N = \sum_{i=1}^{r} N_i$  and  $N_1, N_2, ..., N_r$  are independent, each having a geometric distribution with parameter p.

So  $E[N] = E[\sum_{i=1}^{r} N_i] = \sum_{i=1}^{r} E[N_i] = \sum_{i=1}^{r} \frac{1}{p} = \frac{r}{p}$ 

### Q54

Let X and Y each take on either the value 1 or -1. Let

$$p(1,1) = P\{X = 1, Y = 1\},$$
  

$$p(1,-1) = P\{X = 1, Y = -1\},$$
  

$$p(-1,1) = P\{X = -1, Y = 1\},$$
  

$$p(-1,-1) = P\{X = -1, Y = -1\}$$

Suppose that E[X] = E[Y] = 0. Show that (a) p(1,1) = p(-1,-1); (b) p(1,-1) = p(-1,1). Let p = 2p(1,1). Find (c) Var(X); (d) Var(Y); (e) Cov(X,Y).

#### Answer

(a) 
$$E[X] = E[Y] = 0$$
, so  $0 = E[X] = 1 \cdot p(1,1) + 1 \cdot p(1,-1) - 1 \cdot p(-1,1) - 1$   
 $p(-1,-1) = p(1,1) + p(1,-1) - p(-1,1) - p(-1,-1)$ .  
 $0 = E[Y] = 1 \cdot p(1,1) - 1 \cdot p(1,-1) + 1 \cdot p(-1,1) - 1 \cdot p(-1,-1)$   
 $= p(1,1) - p(1,-1) + p(-1,1) - p(-1,-1)$ 

$$0 = E[X] + E[Y]$$
  
=  $p(1,1) + p(1,-1) - p(-1,1) - p(-1,-1) + p(1,1) - p(1,-1) + p(-1,1)$   
-  $p(-1,-1) = 2p(1,1) - 2p(-1,-1)$ 

which implies  $p(1,1) - p(-1,-1) = 0 \implies p(1,1) = p(-1,-1)$ 

- (b) from (a) we have p(1,1) = p(-1,-1), and 0 = E[X] = p(1,1) + p(1,-1) p(-1,1) p(-1,-1) = p(1,-1) p(-1,1), so we have  $p(1,-1) p(-1,1) = 0 \Rightarrow p(1,-1) = p(-1,1)$ .
- (c)  $Var(X) = E[X^2] (E[X])^2 = E[X^2] = 1[p(1,1) + p(1,-1) + p(-1,1) + p(-1,-1)] = 1$  (No matter X = 1 or  $X = -1, X^2 = 1$ ).
- (d)  $Var(Y) = E[Y^2] (E[Y])^2 = E[Y^2] = 1[p(1,1) + p(1,-1) + p(-1,1) + p(-1,-1)] = 1$  (No matter Y = 1 or Y = -1,  $Y^2 = 1$ ).
- (e)  $Cov(X,Y) = E[XY] E[X] \cdot E[Y] = E[XY] = 1 \cdot p(1,1) 1 \cdot p(1,-1) 1 \cdot p(-1,1) + 1 \cdot p(-1,-1)$ . from (a) and (b) we know that p(1,1) = p(-1,-1) and p(1,-1) = p(-1,1). p = 2p(1,1), so  $p(1,-1) = p(-1,1) = \frac{1}{2} \frac{p}{2}$ . So  $Cov(X,Y) = \frac{p}{2} (\frac{1}{2} \frac{p}{2}) (\frac{1}{2} \frac{p}{2}) + \frac{p}{2} = 2p 1$ .

## Q73

For the multinomial distribution (Exercise 17), let  $N_i$  denote the number of times outcome *i* occurs. Find

(a)  $E[N_i];$ 

(b)  $Var(N_i);$ 

- (c)  $Cov(N_i, N_j);$
- (d) Compute the expected number of outcomes that do not occur.

### Answer

Let's assume that the total number of experiments is n, in each one of the experiment, the outcome i occurs with probability  $p_i$ , so we can see that  $N_i$  has a binomial distribution with parameter n and  $p_i$ .

- (a)  $E[N_i] = n \times p_i = np_i$
- (b)  $Var(N_i) = np_i(1 p_i)$

Let  $X_l = 1$  if in the l - th experiment the outcome *i* occured, otherwise  $X_l = 0$ , for l = 1,2,3, ..., n.  $Y_m = 1$  if in the m - th experiment the outcome *j* occured, otherwise  $Y_m = 0$ , for m = 1,2,3, ..., n

So 
$$N_i = \sum_{l=1}^n X_l$$
 and  $N_j = \sum_{m=1}^n Y_m$ .

(c)  $Cov(N_i, N_j) = Cov(\sum_{l=1}^n X_l, \sum_{m=1}^n Y_m) = \sum_{l=1}^n \sum_{m=1}^n Cov(X_l, Y_m)$ . Based on the condition, we know that if  $l \neq m$  then  $Cov(X_l, Y_m) = 0$ , so  $Cov(N_i, N_j) = \sum_{l=1}^n Cov(X_l, Y_l)$ .

 $Cov(X_l, Y_l) = E[X_lY_l] - E[X_l]E[Y_l]$ , and  $X_lY_l = 0$  (they can't happen at the same time in an experiment).  $E[X_l] = p_i$ ,  $E[Y_l] = p_j$ . Now we can see that  $Cov(X_l, Y_l) = E[X_lY_l] - E[X_l]E[Y_l] = 0 - p_ip_j = -p_ip_j$ , so  $Cov(N_i, N_j) = \sum_{l=1}^n C ov(X_l, Y_l) = n * (-p_ip_j) = -np_ip_j$ .

(d) Let

$$I(N_i = 0) = \begin{cases} 1, \text{ if } N_i = 0\\ 0, \text{ otherwise} \end{cases}$$

So the number of outcomes that do not occur is equal to  $\sum_{i=1}^{r} I(N_i = 0)$ , then  $E[\sum_{i=1}^{r} I(N_i = 0)] = \sum_{i=1}^{r} E[I(N_i = 0)]$ , and we know that  $P(N_i = 0) = (1 - p_i)^n$ , so  $E[I(N_i = 0)] = (1 - p_i)^n$ , now we can see that the expected number of outcomes that do not occur =  $\sum_{i=1}^{r} E[I(N_i = 0)] = \sum_{i=1}^{r} (1 - p_i)^n$